

Bayesian Copula-based Latent Variable Models

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Copulas: The Joys

- ▶ Copulas are mathematical devices used to **model dependence between random variables** regardless of their marginals.
- ▶ Copulas are useful for **data fusion/integration** because they lead to coherent joint models, even when the marginals are in different families (e.g., Gaussian, Poisson, Student, etc) or of different types (e.g, discrete, continuous).
- ▶ Copulas **unlock information contained in the dependence part of the distribution** (second-order) that complements the information in the marginals.
- ▶ Simply put, copulas allow us to **extend statistical methods beyond the use of a multivariate Gaussian or Student**.

At the root of it all, a theorem

- ▶ If Y_1, Y_2, \dots, Y_K are continuous r.v.'s with cdfs F_1, F_2, \dots, F_K , there is an **unique copula** $C : [0, 1]^K \rightarrow [0, 1]$ that links the joint cdf with the marginal ones (Sklar's Theorem).
- ▶ The copula (when $K = 2$) $C : [0, 1] \times [0, 1] \rightarrow [0, 1]$ satisfies

$$F_{12}(t, s) = \Pr(Y_1 \leq t, Y_2 \leq s) = C(F_1(t), F_2(s)).$$

- ▶ The conditional copula satisfies

$$F_{12|X}(t, s) = \Pr(Y_1 \leq t, Y_2 \leq s | X) = C(F_{1|X}(t), F_{2|X}(s) | X)$$

- ▶ Usually we use parametric families so $C(u, v) = C_\theta(u, v)$ such as Clayton's family: $C_\theta(u, v) = [\max(u^{-\theta} + v^{-\theta} - 1, 0)]^{-1/\theta}$.
Frank's family: $K_\theta(u, v) = -\frac{1}{\theta} \ln \left[1 + \frac{(e^{-\theta u} - 1)(e^{-\theta v} - 1)}{e^{-\theta} - 1} \right]$.
- ▶ In a conditional copula, **θ may depend on X** .

Latent Variables (LV)

- ▶ The variable of interest W is sometimes impossible to measure directly
 - ▶ State of the economy
 - ▶ Traffic in a city
 - ▶ State of your health
 - ▶ State of a complex disease
- ▶ Instead, one measures
 - ▶ $\mathbf{Y} = (Y_1, \dots, Y_k)^T$ whose components are surrogates of W and each provide partial information about W
 - ▶ Covariate $\mathbf{X} \in \mathbb{R}^p$
- ▶ We are often interested in the explanatory power of \mathbf{X} for W .

An example

- ▶ Cardiotocography (CTG) is a medical procedure that monitors the fetal heart rate.
- ▶ The LV is the fetus' underlying state of health during birth, W .
- ▶ Our surrogate response is the bivariate vector (Q, Y) where
 - ▶ Q is the number of peaks (acceleration followed by a deceleration of heart beats) for the signal recorded by the CTG
 - ▶ Y is the log of mean short-term "beat-to-beat" variability (MSTV) where the short-term variability (STV) is obtained by measuring the time between successive R waves (cardiac systoles) of the fetus' electrocardiogram.
- ▶ The covariates are FM (fetal movement) and UC (uterine contraction), two continuous variables monitored during birth.

Conditional independence LV model

- ▶ A canonical LV model, given $W_i = X_i\beta + \epsilon$, is

$$Y_i \perp Q_i | W_i$$

$$Y_i \sim N(\mu_c + \lambda_c W_i, \sigma^2)$$

$$Q_i \sim \text{Poisson}(\exp(\mu_d + \lambda_d W_i))$$

- ▶ This implies that the two marginal regressions share a common random effect so they are marginally dependent (and conditionally independent)
- ▶ The induced dependence is not analytically available.

Conditional independence is a Copula LV

- ▶ The copula alternative is, conditional on W_i ,

$$H(Y_i, Q_i|W_i) = C_{\theta_i}(F_Y(Y_i|W_i), F_Q(Q_i|W_i)), \quad \theta_i = \kappa^{-1}(\xi_0 + \xi_1 W_i)$$
$$Y_i \sim N(\mu_c + \lambda_c W_i, \sigma^2); \quad Q_i \sim \text{Poisson}(\exp(\mu_d + \lambda_d W_i))$$

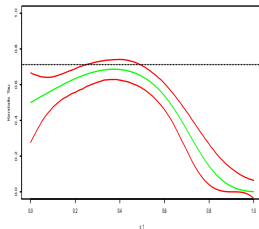
- ▶ The whole joint distribution of (Y, Q) is varying with W not just the marginals.
- ▶ The copula captures the residual dependence on W after the marginal effects have been accounted for.
- ▶ The previous model is obtained when the copula is the independence copula.

Why the Conditional Copula?

- ▶ $Y_i|x \sim N(f_i(x), \sigma_i) \quad x \in \mathbb{R}^2$
- ▶ True marginal means:
 - ▶ $f_1(x) = 0.6 \sin(5x_1) - 0.9 \sin(2x_2)$
 - ▶ $f_2(x) = 0.6 \sin(3x_1 + 5x_2)$
 - ▶ $\sigma_1 = \sigma_2 = 0.2, \textcolor{red}{X_1} \perp \textcolor{red}{X_2}.$
- ▶ Copula (Levi and Craiu, 2018): $\theta(x) = 0.71$
- ▶ $\textcolor{red}{\text{Suppose } x_2 \text{ is not observed so inference is based only on } x_1}$

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CTG: The LV Copula Model

- ▶ $(Q_i, Y_i)|W_i$ has joint density

$$f_{(Q,Y)}(q,y) = f_c(y) \cdot [C_{d|c}(F_d(q), F_c(y)) - C_{d|c}(F_d(q-), F_c(y))],$$

where

$$C_{d|c}(u_d, u_c) = \frac{\partial}{\partial u_c} C(u_d, u_c).$$

- ▶ Data Augmentation: Introduce latent variable Z such that

$$Q \stackrel{d}{=} F_d^-(F_Z(Z)),$$

- ▶ The copula between (Y, Z) is the same as the copula between (Y, Q)
- ▶ We can choose the distribution of Z to help the computation.
- ▶ For instance if we use a Gaussian copula, it helps to have $Z \sim N(0, 1)$
- ▶ [Craiu and Sabeti \(2012\)](#); [Smith and Khaled \(2012\)](#).

The Augmented LV Copula Model for the CTG Example

- The augmented model for CTG data is

$$Z_i \sim \mathcal{N}(0, 1)$$

$$Q_i \mid W_i \sim \text{Poisson}(e^{\mu_d + \lambda_d W_i})$$

$$Y_i \mid W_i \sim \mathcal{N}(\mu_c + \lambda_c W_i, \sigma^2)$$

$$(Z_i, Y_i) \mid W_i \sim C^{\text{Gauss}} \left(\Phi(\cdot), \Phi \left(\frac{\cdot - \mu_c - \lambda_c W_i}{\sigma} \right) \mid \theta(W_i, \xi) \right)$$

$$W_i \sim \mathcal{N}(x_i \beta, 1),$$

CTG: The Augmented LV Copula Model

- ▶ The dependence between Y , Z and Q is defined by their joint conditional distribution

$$f_{(Q,Z,Y)|W}(q, z, y | w) = h(z, y | w, \mu_c, \lambda_c, \psi_c, \xi) \cdot \mathbb{1}_{F_Z^{-1}(F_d(q - |\varphi_d(\mu_d, \lambda_d, w))) \leq z < F_Z^{-1}(F_d(q | \varphi_d(\mu_d, \lambda_d, w)))}.$$

- ▶ Let $\xi = (\xi_0, \xi_1) \in \mathbb{R}^2$ and $A(w) = \xi_0 + \xi_1 \cdot w$. Then we set

$$\theta(w, \xi) = \frac{e^{A(w)} - e^{-A(w)}}{e^{A(w)} - e^{-A(w)}}$$

as the correlation parameter of the bivariate Gaussian conditional copula of $(Y, Z) | W = w$.

- ▶ Parameters are a priori independent

Some MCMC details

- ▶ If the copula and marginals are Gaussian the joint is a multivariate normal so some of the conditional densities are available in closed form.
- ▶ For other copula families we rely on MwG moves.
- ▶ We sample $\{Z_i : 1 \leq i \leq n\}$ from its conditional distribution and use the samples only to update the copula parameters ξ .
- ▶ Good initialization helps:
 - ▶ $\beta^{(0)} = (\mathbf{X}^\top \mathbf{X})^{-1} \mathbf{X}^\top \mathbf{y}$
 - ▶ $W_i^{(0)} = (\beta^{(0)})^\top \mathbf{x}_i$
 - ▶ $(\mu_d^{(0)}, \lambda_d^{(0)})$ is the MLE based on the marginal likelihood, etc
- ▶ Adaptive strategy for all MwG: target an acceptance rate of 44%.

Model Selection: WAIC

- The WAIC is defined as

$$\text{WAIC}(\mathcal{M}) = -2\text{fit}(\mathcal{M}) + 2p(\mathcal{M}), \quad (1)$$

where the model fitness is

$$\text{fit}(\mathcal{M}) = \sum_{i=1}^n \log(\mathbb{E}[\text{Pr}(y_i, q_i | \omega, \mathcal{M})]) \quad (2)$$

and the penalty

$$p(\mathcal{M}) = \sum_{i=1}^n \text{Var}(\log(\text{Pr}(y_i, q_i | \omega, \mathcal{M}))), \quad (3)$$

where ω contains all the parameters and latent variables in the model.

Spotlight on dependence: A conditional WAIC

- We use the following two conditional WAICs (Levi and Craiu, 2018)

$$\begin{aligned} \text{CWAIC}_{Y|Q}(\mathcal{M}) &= -2 \sum_{i=1}^n \log(\mathbb{E}[\Pr(y_i|q_i, \omega, \mathcal{M})]) + \\ &\quad + 2 \sum_{i=1}^n \text{Var}(\log(\Pr(y_i|q_i, \omega, \mathcal{M}))), \\ \text{CWAIC}_{Q|Y}(\mathcal{M}) &= -2 \sum_{i=1}^n \log(\mathbb{E}[\Pr(q_i|y_i, \omega, \mathcal{M})]) + \\ &\quad + 2 \sum_{i=1}^n \text{Var}(\log(\Pr(q_i|y_i, \omega, \mathcal{M}))), \end{aligned}$$

- $\frac{1}{2}(\text{CWAIC}_{1|2} + \text{CWAIC}_{2|1})$ is asymptotically equivalent to CCV for the marginal likelihood

$$\text{CCV}(\mathcal{M}) = \frac{1}{2} \left\{ \sum_{i=1}^n \log(\Pr(y_i|q_i, \mathcal{D}_{-i}, \mathcal{M})) + \sum_{i=1}^n \log(\Pr(q_i|y_i, \mathcal{D}_{-i}, \mathcal{M})) \right\}.$$

Simulation Experiment

- Generate data using a Gaussian copula

Gaussian copula

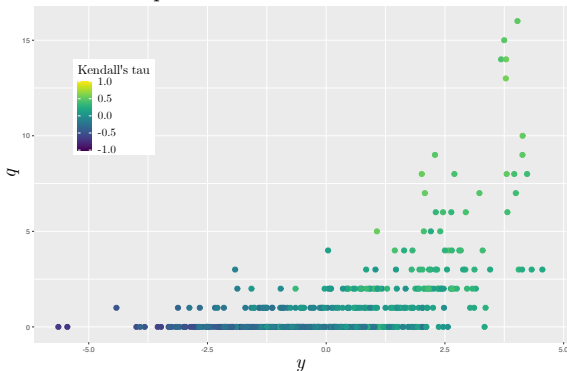


Figure: Bivariate scatterplot of the generated data with Gaussian copula, and Poisson and normal marginals

Simulation Experiment

- $\text{CWAIC}_{Y|Q}$ and $\text{CWAIC}_{Q|Y}$ selection criteria

Criteria\Copula	Gaussian	Frank	Gumbel	Clayton	Indep
$\text{CWAIC}_{Y Q}$	1627.36	1642.36	2395.17	1637.17	1606.31
$\text{CWAIC}_{Q Y}$	950.71	982.42	1673.57	976.05	997.43
Average	1289.04	1312.39	2034.37	1306.61	1301.87

Simulation Experiment

Gaussian copula

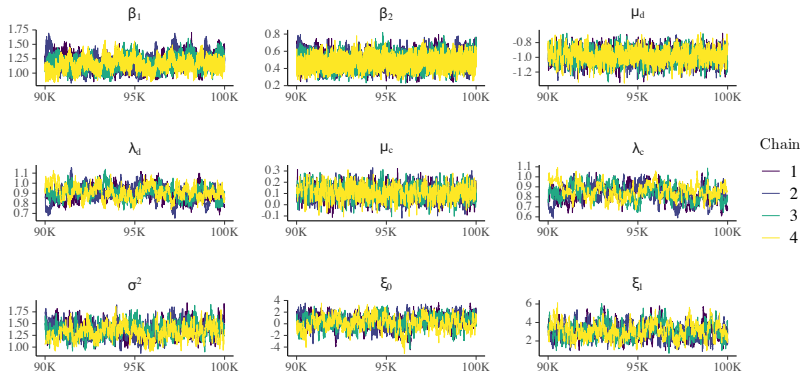
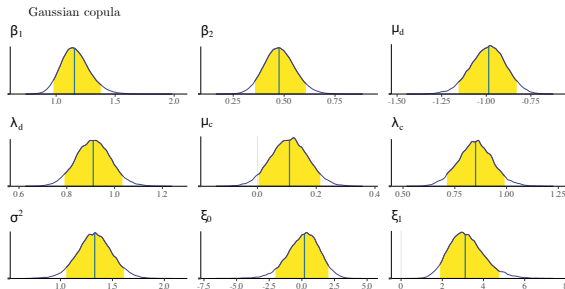


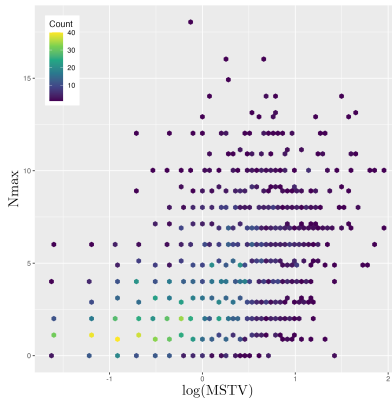
Figure: Traceplots for η 's components.

Simulation Experiment



	β_1	β_2	λ_d	λ_c	ξ_1
Mean	1.18	0.48	0.90	0.84	3.10
True	1	0.5	1	1	3

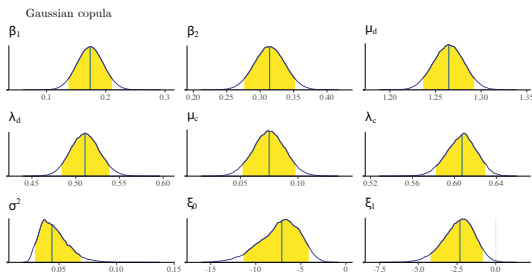
CTG: The data



CTG: Estimates

- ▶ $WAIC$, $WAIC_{Y|Q}$ and $WAIC_{Q|Y}$ all point to the Gaussian copula (over Gumbel, Frank, Clayton, Independence).
- ▶ The posterior means

	β_1 (FM)	β_2 (UC)	λ_d	λ_c	ξ_1
Mean	0.1744	0.3147	0.5101	0.6038	-2.3401



The Past & Future

- ▶ Copulas offer a way to bypass the paucity of available joint distributions.
- ▶ Copulas allow the integration of multiple (dependent) sources of information/data via joint modeling
- ▶ Joint models can be used for prediction/imputation of an expensive variable given values for cheaper ones.
- ▶ So far have been used to further empower multivariate regressions, time series, HMMs, LV models, etc
- ▶ Computational challenges, especially in higher dimensions
- ▶ Papers available here:
<https://www.utstat.toronto.edu/craiu/Papers/index.html>

References

- Craiu, R. V. and Sabeti, A. (2012). In mixed company: Bayesian inference for bivariate conditional copula models with discrete and continuous outcomes. J. Multivariate Anal., 110:106–120.
- Levi, E. and Craiu, R. V. (2018). Bayesian inference for conditional copulas using Gaussian process single index models. Computational Statistics & Data Analysis, 122:115–134.
- Smith, M. S. and Khaled, M. A. (2012). Estimation of copula models with discrete margins via bayesian data augmentation. Journal of the American Statistical Association, 107(497):290–303.